Phase angle from bioelectrical impedance analysis remains an independent predictive marker in HIV-infected patients in the era of highly active antiretroviral treatment\(^1,2\)

Achim Schwenk, Alexander Beisenherz, Katja Römer, Gisela Kremer, Bernd Salzberger, and Marinos Elia

**ABSTRACT**

Background: Highly active antiretroviral treatment (HAART) reduces the risk of wasting in HIV infection and may alter the prognostic weight of wasting. The phase angle from bioelectrical impedance analysis (BIA) can be interpreted as a surrogate marker for the catabolic reaction to chronic HIV infection and opportunistic disease.

Objective: Our objective was to assess the prognostic ability of the phase angle in HIV-infected patients in the era of HAART.

Design: Two cross-sectional observation studies were conducted in 1996 and 1997 at a German university outpatient HIV clinic. In the 1996 and 1997 cohorts, HAART was prescribed to 17 of 212 and 168 of 257 patients at baseline and to 179 of 212 and 234 of 257 patients during observation, respectively. Whole-body BIA was assessed at 50 KHz. Time to clinical progression and survival were calculated by using Cox proportional hazard models with time-dependent covariates. Median observation times were 1000 and 515 d for the 1996 and 1997 cohorts, respectively.

Results: Higher phase angle was associated with a lower relative mortality risk, adjusted for viral load and CD4\(^+\) cell count, of 0.49 (95% CI: 0.30, 0.81) per degree in 1996 and of 0.33 (95% CI: 0.18, 0.61) in 1997. The influence of phase angle on time to clinical progression, adjusted for viral load and CD4\(^+\) cell count, was not significant in 1996 but the relative risk was 0.58 (0.36, 0.83) in 1997.


**KEY WORDS** Acquired immunodeficiency syndrome, AIDS, biological markers, body fluid compartments, body composition, bioelectrical impedance, HIV infection, survival analysis

**INTRODUCTION**

Loss of body weight and wasting of lean body mass are leading symptoms of the advanced stages of HIV infection. Early in the epidemic it was proposed that malnutrition may be an important cofactor of disease progression (1). Several studies undertaken before the era of highly active antiretroviral treatment (HAART) showed that weight loss was associated with more rapid disease progression and shorter survival after CD4\(^+\) cell count was controlled for (2, 3). Body-composition studies done with bioelectrical impedance analysis (BIA) suggested that low body cell mass was an adverse prognostic marker (4). The correlation of BIA and prognosis was found to be particularly strong when a measured BIA parameter, phase angle, was used instead of the derived body-composition estimates (5). However, changes in weight and fat-free mass (FFM) had no significant influence on disease progression of asymptomatic HIV-infected patients in a study that measured body composition with another method (6). Since these studies were conducted, the prognosis of HIV-infected patients in industrialized countries has changed considerably. HAART has reduced morbidity and mortality (7, 8), including the incidence of malnutrition (9). Additionally, viral load has been found to be a much more powerful predictor of risk than is CD4\(^+\) cell count (10).

The current study was undertaken to determine whether a low phase angle with BIA is still associated with a poor prognosis if considered together with other state-of-the-art prognostic markers and their changes during antiretroviral treatment. The phase angle is defined as the relation between the 2 vector components of impedance: resistance and reactance. It may be interpreted as an indicator of water distribution between the extra- and intra-cellular spaces. However, the relation of impedance to body composition is indirect and incompletely understood (11). Therefore, the phase angle, rather than derived body-composition estimates, was examined for its prognostic weight in the current study.

This study was part of a series of cross-sectional and longitudinal BIA measurements in a defined population of HIV-infected outpatients in the era of rapid improvements in antiretroviral therapy. Data on the effect of protease inhibitors on weight and BIA data have been published elsewhere (9, 12). In this article, we report the prognostic weight of phase angle for survival and progression-free survival in the same population, before and after the introduction of HAART.

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BIA PHASE ANGLE AS PREDICTIVE MARKER IN HIV

TABLE 1
Baseline characteristics and outcomes

<table>
<thead>
<tr>
<th></th>
<th>1996 Cohort (n = 181 M, 31 F)</th>
<th>1997 Cohort (n = 220 M, 37 F)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (y)</td>
<td>40.4 ± 10.6(^{2})</td>
<td>39.7 ± 10.5</td>
</tr>
<tr>
<td>BMI (kg/m(^{2}))</td>
<td>22.6 ± 3.1</td>
<td>23.1 ± 3.3(^{2})</td>
</tr>
<tr>
<td>Total weight change</td>
<td>-3.1 ± 10.3</td>
<td>0.9 ± 9.8(^{2})</td>
</tr>
<tr>
<td>(% of usual body weight)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>CD4(^{+}) cell count (\times 10^{3}) cells/L</td>
<td>203 ± 209</td>
<td>306 ± 212(^{*})</td>
</tr>
<tr>
<td>HIV RNA (\log_{10}) copies/L</td>
<td>7.3 ± 1.1</td>
<td>6.3 ± 1.3(^{11})</td>
</tr>
<tr>
<td>HAART at baseline(^{2})</td>
<td>17 (8)(^{4})</td>
<td>168 (65)(^{4})</td>
</tr>
<tr>
<td>HAART during observation time</td>
<td>179 (84)</td>
<td>234 (91)</td>
</tr>
</tbody>
</table>

Outcomes

- No progression: 153 (72) vs. 233 (91)\(^{4}\)
- Progression, survived: 32 (15) vs. 14 (5)
- Death, no other progression: 13 (6) vs. 4 (2)
- Death after progression: 14 (7) vs. 6 (2)
- Observation time (d): 1000 (28–1203)\(^{7}\) vs. 515 (28–704)\(^{7}\)

\(^{1}\)132 of 979 cohort were also in 1996 cohort.
\(^{2}\)\(\bar{x} \pm SD\).
\(^{3}\)Significantly different from 1996 cohort: \(\bar{x} \pm 0.05, \bar{x} \pm 0.01\). HAART, highly active antiretroviral treatment.
\(^{4}\)n; percentage in parentheses.
\(^{7}\)Median; range in parentheses.

SUBJECTS AND METHODS

The study was conducted at the outpatient clinic for infectious diseases of a large university hospital in western Germany. All HIV antibody–positive outpatients attending the department during 2 periods of 6 wk each were asked to participate. The 2 periods in March–April 1996 and July–August 1997 were 6 wk each were asked to participate. The 2 periods in March–April 1996 and July–August 1997 were 6 wk apart. Patients were excluded from the analysis if they were receiving enteral or parenteral nutrition or any drug treatment for wasting, lacked documentation of HIV viremia or CD4\(^{+}\) cell count, or had had follow-up of < 2 wk, unless progression or death occurred during this period.

During the period in 1996, 250 of 254 clinic patients consented to BIA measurement. The current analysis was restricted to 212 patients because 3 patients received parenteral nutrition, 20 had an insufficient length of follow-up, and 15 had missing data. During the period in 1997, 266 of 276 patients participated but 9 patients were excluded because their follow-up periods were too short. Baseline characteristics of the remaining 212 and 257 patients are given in Table 1.

Whole-body bioelectrical impedance was measured by using a BIA 2000–1 device (Data Input, Frankfurt, Germany) (13). The relation of the 2 impedance components at 50 KHz, reactance \(X_{c}\) and resistance \(R\), was expressed as phase angle \(\alpha = (X_{c} \times 180/\pi) / (R \times \pi)\). Estimates for intracellular water (ICW), total body water (TBW), and FFM were calculated by using regression equations resulting from the comparison of BIA with reference methods in HIV-infected and healthy subjects (14). Extracellular water was calculated by subtraction as ECW = TBW – ICW, and fat mass as FM = body weight – FFM. All measurements were made by the same investigator (AB) using standard electrode positions (15). HIV viral load was assessed at the Institute of Virology, University of Cologne, by using quantitative polymerase chain reaction (PCR; Roche Amplicor; Hoffmann-La Roche, Germany). Detection limits for this test were 300 copies/mL before March 1997 and 20 copies/mL afterward. All results with undetectable copy numbers were set to these detection limits. Results obtained by the ultrasensitive PCR were retested with the less sensitive assay if they were above the linear range of the assay (> 10000 copies). Routine methods were used for all other laboratory tests.

Statistics were calculated by using SPSS version 7.5 (16). Comparisons were made with Pearson’s chi-square for dichotomous variables and Student’s t test for continuous variables. The prognostic influence of BIA and other independent variables on survival and time to clinical progression were tested with the Kaplan-Meier method for monovariate analysis. Proportional hazards were calculated in mono- and multivariate Cox models. Viral load and CD4\(^{+}\) cell counts were used as time-dependent covariates by using their mean \(\log_{10}\)-transformed results within 100-d periods after the periods in 1996 and 1997, respectively (16). Variables in the multivariate models were selected with the backward elimination log-likelihood method and with thresholds of \(\alpha < 0.05\) for entry and \(\alpha \geq 0.10\) for removal of variables. Phase angle was the principal BIA-derived independent variable in these models but was replaced by the 3 body compartments—body fat, ICW, and ECW—or by the ECW-ICW ratio in other explorative models. Time to clinical progression was defined as the time between baseline and the first episode of any disease listed in the Centers for Disease Control and Prevention case definition for AIDS (category C) (17), or death. Recurrence of diseases that had occurred before the period in 1996 were counted as new episodes if the period without symptoms or specific induction treatment was ≥ 1 mo. Interactions between the phase angle and other independent variables were further explored by using general linear models (16). The 1996 cohort was observed for 1000 (28–1203) d [median (range)], and the 1997 cohort for 515 (28–704) d. The 2 cohorts overlapped with 132 patients, but only 3 deaths and 1 clinical progression event were counted twice. Omission of these events from the 1996 cohort data did not alter the overall results (data not shown). The study was conducted in accordance with the Declaration of Helsinki.

RESULTS

A greater proportion of patients in the 1996 cohort than in the 1997 cohort developed an AIDS-defining event or died, as shown in Table 1. Likewise, BIA results suggested healthier body compositions in the 1997 cohort than in the 1996 cohort, as reported previously (9). In both cohorts, a lower phase angle with BIA was significantly associated with shorter survival and time to clinical progression, as shown in Figure 1. Survival after 1996 was significantly shorter for patients in the lowest quartile of the phase angle (< 5.3\(^{\circ}\): 878 d [95% CI: 758, 998] compared with 1013 d (985, 1041) in the 3 higher quartiles together (\(P < 0.001\)). When the same thresholds were applied to a similar analysis in the 1997 cohort, patients with a phase angle < 5.3\(^{\circ}\) again had a shorter estimated survival of 463 d (397, 528) compared with 697 d (690, 705) for patients with a higher phase angle (\(P < 0.001\)). All deaths in both cohorts were either directly attributable to opportunistic infections or tumors or occurred in advanced AIDS cases without other apparent reasons. The influence of phase angle on time to next clinical progression was not significant in the 1996 cohort. In the 1997 cohort, patients with a phase angle < 5.3\(^{\circ}\) had a clinical progression after 406 d (330, 483) compared with 670 d (652, 688) in patients with a higher phase angle (\(P < 0.001\)). In monovariate Cox models with survival as the endpoint, the relative risk (RR) reduction associated with an increase of the
phase angle by 1° was 0.33 (95% CI: 0.21, 0.52) in the 1996 cohort and 0.29 (0.16, 0.53) in the 1997 cohort. In comparison, a log_{10} increase of CD4^+ cells was associated with an RR of 0.19 (0.12, 0.32) in 1996 and 0.11 (0.06, 0.22) in 1997. A log_{10} increase of viral load was associated with an RR of 2.64 (1.76, 3.97) in 1996, but in 1997, this association was less pronounced with an RR of 1.84 (1.14, 2.67). Patients with prior AIDS had a 7.56 (2.61, 21.9) times higher mortality risk in 1996, but this association was only marginally significant in 1997 (RR: 3.82; 0.99, 14.8). Age, sex, and body mass index had no significant influence on the relative mortality risk. Clinical progression, rather than death, was the endpoint in another set of similar monovariate models. Again, significant associations with RR were seen with the same independent variables, and body mass index was not associated with the risk of progression.

In multivariate Cox models, survival time after 1996 was predicted only by the CD4^+ cell count, HIV viral load, and the phase angle (Table 2). Survival after 1997 was predicted only by phase angle and CD4^+ cell count, not by viral load. Time to clinical progression was predicted by the CD4^+ cell count in both cohorts, together with a prior AIDS diagnosis in 1996, and together with the phase angle in 1997 (Table 2). Viral load did not contribute significantly to the prediction of clinical progression in either cohort.

Significant correlations between CD4^+ cell count, viral load, prior AIDS, and the phase angle were observed. They were further explored in general linear models (Table 3). Together with age, sex, and body mass index, these variables explained 33.8% and 25.6% of the variance in phase angle in 1996 and 1997, respectively. However, the contribution of CD4^+ cell count, viral load, and prior AIDS to this model decreased from 17.5% in 1996 to 5.6% in 1997. The roles of these 3 variables also differed between the cohorts. In 1996, prior AIDS was the most powerful predictor of a low phase angle, and high viral load had a moderate influence. In 1997, neither of these variables was significantly associated with the phase angle, whereas a lower CD4^+ cell count was associated with a lower phase angle (Table 3).
As expected from theory, phase angle was strongly correlated with the BIA estimate of the ECW-ICW ratio \((r = 0.82\) in men, \(r = 0.65\) in women, \(P < 0.001\)). An increase of phase angle by \(1^\circ\) corresponded to a decrease of the ratio by the factor 0.901 (0.897, 0.906).

### DISCUSSION

The results of this study show a strong ability of the phase angle to predict survival and clinical progression in HIV-infected patients, independent of the degree of immunodeficiency and viremia. The introduction of HAART to this population has not eliminated the prognostic role of the phase angle shown earlier in the HIV epidemic (5). However, the percentage of patients with a low phase angle (arbitrarily defined as the lowest quartile of the 1996 cohort: \(< 5.3^\circ\)) decreased from 26% to 12%, as shown in detail elsewhere (9). Likewise, the risk of clinical progression and death was reduced considerably (Table 1).

What pathogenic effects are represented in the phase angle? Although BIA is an established technique for assessment of body composition in health and disease (11), the relation of measured impedance to body composition is indirect and not fully clarified. Phase angle describes the relation between the 2 vector components of impedance (reactance and resistance) of the human body to an alternating electric current. Because the current passes only through the ionized water compartments within the body, the volume of TBW can be estimated from resistance (11). Reactance reflects the ability of cell membranes to act as imperfect capacitors. Therefore, phase angle is an indicator of the distribution of water between the intra- and extracellular spaces (18). A high phase angle corresponds to a low ECW-ICW ratio.

In the current study, the prognostic power of the phase angle was much stronger than that of any of the 3 body compartments estimated from BIA. In HIV-negative patients with bacteremia, our study group also found the ratio of 2 other impedance parameters [resistance at zero \((R_0)\) and resistance at infinite frequency of alternate current \((R_{\infty})\)] to be superior to the calculated ECW-ICW ratio as a marker of adverse prognosis (19). These findings indicate that it may be easier to define normal and pathologic ranges for crude BIA data than for derived estimates of body composition. More importantly, bioelectrical impedance may reflect pathogenic events beyond its correlation with body composition. Expansion of ECW and loss of ICW are typical features of systemic illness (20–22) but unlikely candidates for a direct pathogenic effect. Rather, they accompany protein leakage into the extracellular space and loss of intracellular protein in critical illness (23). Tumor necrosis factor \(\alpha\) elicits capillary leakage in animal models (24), partly mediated by nitric oxide (25).

Other surrogate markers for these pathogenic events are more widely known to predict survival than is the phase angle. Loss of intracellular potassium and extracellular accumulation of sodium result in an increased whole-body exchangeable Na\(^+-\)K\(^+\) ratio, which is a strong predictor of mortality in surgical patients (26). Hypoaalbuminemia results largely from protein leakage (23) and is an adverse prognostic marker in systemic illness, such as bacteremia (27), HIV infection (28), and tuberculosis (29). Like the Na\(^+-\)K\(^+\) ratio and serum albumin, phase angle can be interpreted as a global marker of the systemic reaction that forms an integral part of the host defense to systemic infection but may eventually result in malnutrition (30). Further studies comparing the phase angle with more direct assessment of metabolic distress will be needed to test this hypothesis.

If one accepts this interpretation of the phase angle, our findings indicate that the systemic response that predisposes patients

### TABLE 2

Multivariate Cox proportional hazard models

<table>
<thead>
<tr>
<th>Independent variable</th>
<th>1996 Cohort ((n = 212))</th>
<th>1997 Cohort ((n = 257))</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Endpoint: death</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Phase angle (^1)</td>
<td>0.49 (^a)</td>
<td>0.33 (^a)</td>
</tr>
<tr>
<td>CD4 (^+) cell count (^3)</td>
<td>0.40 (^a)</td>
<td>0.10 (^a)</td>
</tr>
<tr>
<td>HIV viremia (^2)</td>
<td>1.82 (^a)</td>
<td>(1.16, 2.84)</td>
</tr>
<tr>
<td><strong>Endpoint: clinical progression</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Phase angle (^1)</td>
<td>0.99 (^d)</td>
<td>0.58 (^d)</td>
</tr>
<tr>
<td>CD4 (^+) cell count (^3)</td>
<td>0.31 (^d)</td>
<td>0.15 (^d)</td>
</tr>
<tr>
<td>Prior AIDS (^1)</td>
<td>4.20 (^d)</td>
<td>(1.24) (^d)</td>
</tr>
</tbody>
</table>

\(^1\)For observation times, see Table 1.

\(^2\)Relative risk (Cox proportional hazard).

\(^3\)CD4 \(^+\) cell count and HIV viral load were treated as time-dependent covariates.

\(^a\)For variables without significant contribution to the final model, RRs in parentheses are derived from the forced addition of this variable to the final model.

### TABLE 3

Determinants of the phase angle

<table>
<thead>
<tr>
<th>Determinant</th>
<th>1996 Cohort parameter estimates (95% CI)</th>
<th>1997 Cohort parameter estimates (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Prior AIDS (^1)</td>
<td>0.69 (^d) (0.84, 0.56)</td>
<td>0.88 (1.08, 0.72)</td>
</tr>
<tr>
<td>Baseline CD4 (^+) count ((\log_{10}))</td>
<td>1.06 (0.91, 1.24)</td>
<td>1.31 (^d) (1.02, 1.69)</td>
</tr>
<tr>
<td>Baseline HIV RNA ((\log_{10}))</td>
<td>0.92 (^d) (0.84, 1.00)</td>
<td>1.00 (0.92, 1.09)</td>
</tr>
<tr>
<td>Male sex (^3)</td>
<td>1.44 (^d) (1.1, 1.87)</td>
<td>1.91 (^d) (1.46, 2.49)</td>
</tr>
<tr>
<td>Age (^3)</td>
<td>0.97 (^d) (0.96, 0.98)</td>
<td>0.97 (^d) (0.96, 0.98)</td>
</tr>
<tr>
<td>BMI ((\text{kg/m}^2))</td>
<td>1.09 (^d) (1.05, 1.12)</td>
<td>1.09 (^d) (1.06, 1.12)</td>
</tr>
</tbody>
</table>

\(^1\)Parameter estimates in general linear models predicting the phase angle. For example, for 2 patients who differed in baseline HIV RNA in 1996 by \(1 \log_{10}\) but were identical in all other variables, the patient with higher HIV RNA would be expected to have a 0.92 times lower phase angle.

\(^2\)P < 0.001.

\(^3\)P < 0.05.
to malnutrition and wasting is an independent risk factor in HIV infection together with viral load and immunodeficiency. Malnutrition, when more strictly defined by a low body mass index, was not significantly associated with prognosis in the current study. Effective antiretroviral treatment reduces the risk of wasting, controls viremia, and restores immune competence. Although statistically independent, these 3 effects are inseparable in patients. Different associations between phase angle and other measures of disease severity were found between 1996 and 1997. In the 1996 cohort, the patients’ nutritional status was largely determined by their history of opportunistic infections, reflecting the episodic nature of malnutrition in HIV infection (31, 32).

Effective antiretroviral treatment not only led to a lower incidence of opportunistic infections before 1997 but may also have reduced the severity of episodes and their metabolic consequences. This may have unmasked the metabolic effect of chronic HIV infection, as reflected in the CD4+ cell count.

Metabolic adverse effects of HAART may have confounded our data. Because this prospective study was designed before the first description of the fat redistribution syndrome, or lipodystrophy (33, 34), incidence of this syndrome could be determined only retrospectively in a subset (n = 111) of this population. As described elsewhere (9), the fat redistribution syndrome was associated with a greater increase of the phase angle between 1996 and 1997. Diagnosis of the fat redistribution syndrome did not have a detectable influence on prognosis in this subset (data not shown), but the statistical power of this finding is small. A protective effect of the syndrome is unlikely, apart from its association with low viral load (34, 35); hence, it could either have no effect or lead to underestimation of the prognostic power of phase angle.

Because of the close association between metabolic status and other manifestations of HIV infection, our data do not provide a causal link between reversal of catabolism and improved prognosis. However, they do underline the importance of monitoring the patient’s metabolic status alongside viral load and CD4+ cell count in assessing prognosis. Phase angle with BIA may become a useful surrogate marker for the systemic response to chronic HIV infection.

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REFERENCES